

Evolution de la diversité génétique d'une population ayant une structure spatiale

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Genetic diversity and spatial structure

Aim: Model and understand the evolution of the genetic diversity of a population having a continuous spatial structure.

↪ Dimension 2 is the most relevant for applications to biological pop., but the mathematical models are interesting in any dimension.



Footprint of a spatial structure

- ▶ Interactions/reproduction require that individuals should be sufficiently close to each other.
- ▶ Offspring are born in a more or less extended neighbourhood of their parents.
- ▶ The selective advantage offered by certain alleles depends on the environment, which can vary from one region to another.

↪ Local allelic distributions are correlated, in a specific way that depends on parameters such as the speed of spatial diffusion of genes across the population.

Questions of interest

- ▶ (Compound) parameters characterising the genetic diversity in a population and its evolution?
- ▶ Correlation pattern between the local allelic distributions at several locations, under different evolutionary scenarios?
Inference methods for the estimation of the corresponding key parameters?
- ▶ Detection of evolutionary forces in action based on appropriate types of data?



The spatial Λ -Fleming-Viot process

(Barton & Etheridge, 2008)

Allele/type space : K compact ($K = \{0, 1\}$, for example).

Population at time t : Measure M_t on $\mathbb{R}^d \times K$ whose first marginal distribution is Lebesgue measure (uniform density of indiv.). Possible decomposition:

$$M_t(dx, dk) = dx \rho_t(x, dk).$$

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Evolution : μ a σ -finite measure on \mathbb{R}_+^* , $\{\nu_r, r > 0\}$ a collection of probability measures on $[0, 1]$. Let Π be a Poisson point process on $\mathbb{R} \times \mathbb{R}^d \times \mathbb{R}_+^* \times [0, 1]$ with intensity measure $dt \otimes dx \otimes \mu(dr)\nu_r(du)$.

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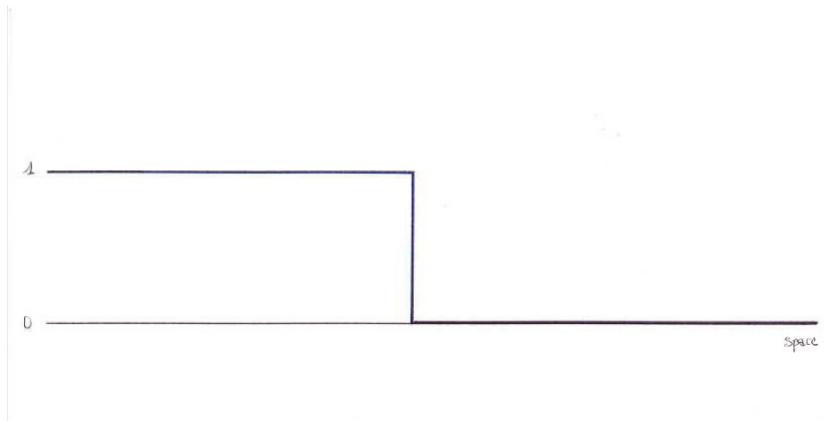
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If $(t, x, r, u) \in \Pi$, at time t and in $B(x, r)$:

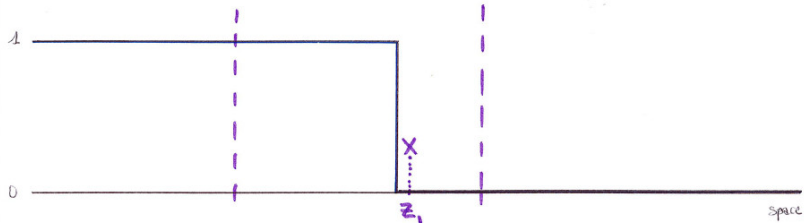
- ▶ We choose a parental allele κ according to the allele distribution at time $t-$ within $B(x, r)$;
- ▶ For every $y \in B(x, r)$, $\rho_t(y, dk) = (1 - u)\rho_{t-}(y, dk) + u\delta_\kappa(dk)$.

An example



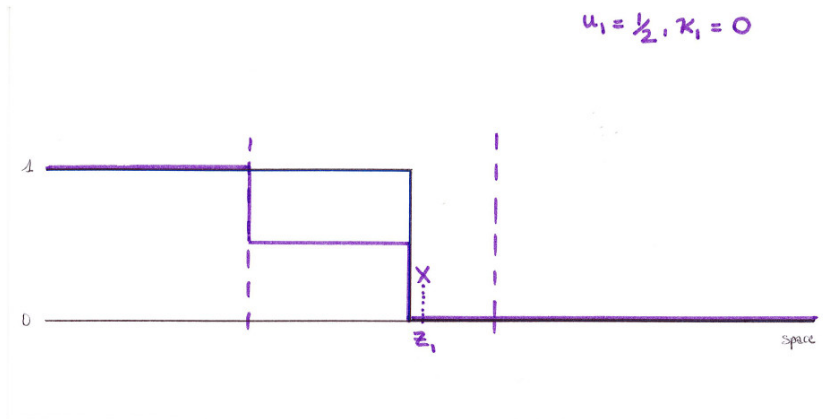
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$$u_1 = \frac{1}{2}, \kappa_1 = 0$$



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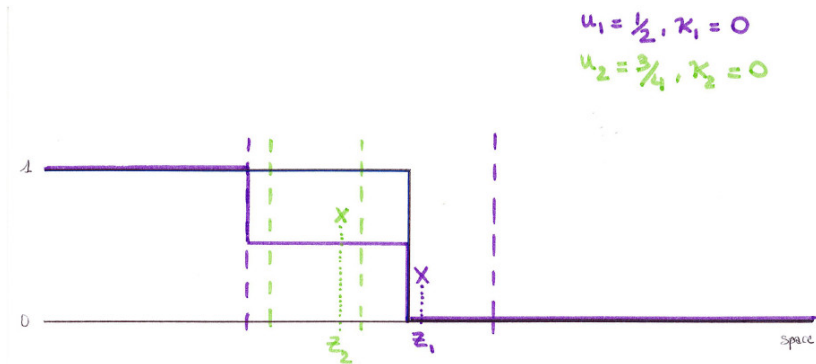
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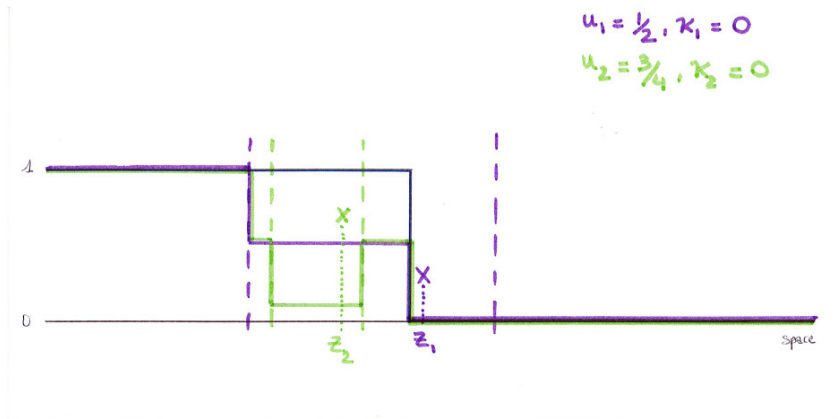
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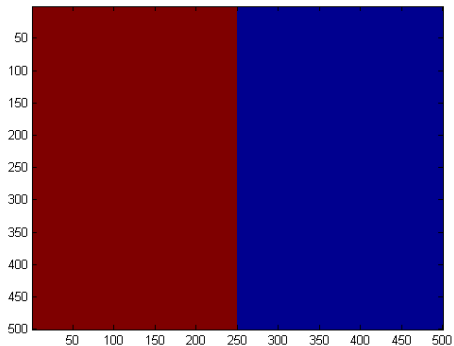
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Over larger scales

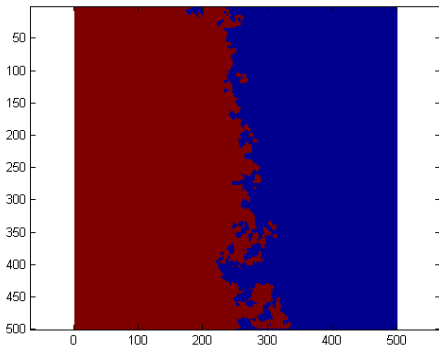
Initial configuration:



Simulations by H. Saadi. Fixed radius, $\nu \equiv 1$.

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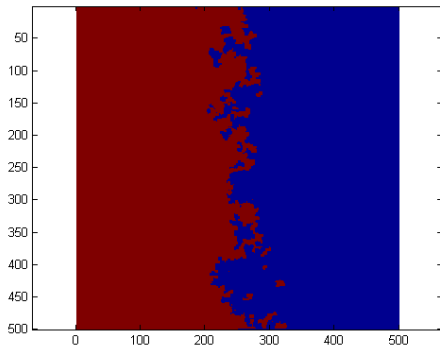
After $2 \cdot 10^6$ events:



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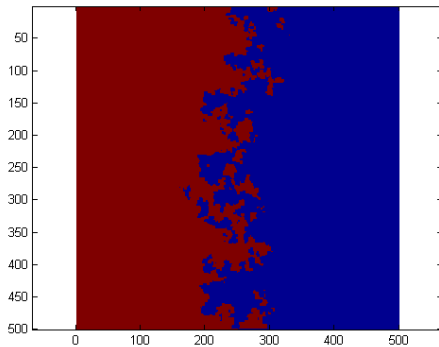
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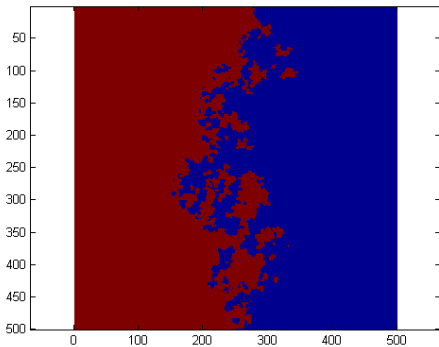
After $4 \cdot 10^6$ events:



Simulations by H. Saadi. Fixed radius, $\nu \equiv 1$.

Over larger scales

After $5 \cdot 10^6$ events:



Simulations by H. Saadi. Fixed radius, $\nu \equiv 1$.

A few remarks

Existence/uniqueness:

- ▶ Initially obtained in Barton et al. (2010) via an analytical method due to Evans (1997),
- ▶ then via a particle construction (V. & Wakolbinger, 2015; Etheridge & Kurtz, “2019”).

Sufficient condition:

$$\int_0^\infty \int_0^1 ur^d \nu_r(du) \mu(dr) < \infty.$$

State space: \mathcal{M}_λ , endowed with the topology of vague convergence (which makes it a compact space).

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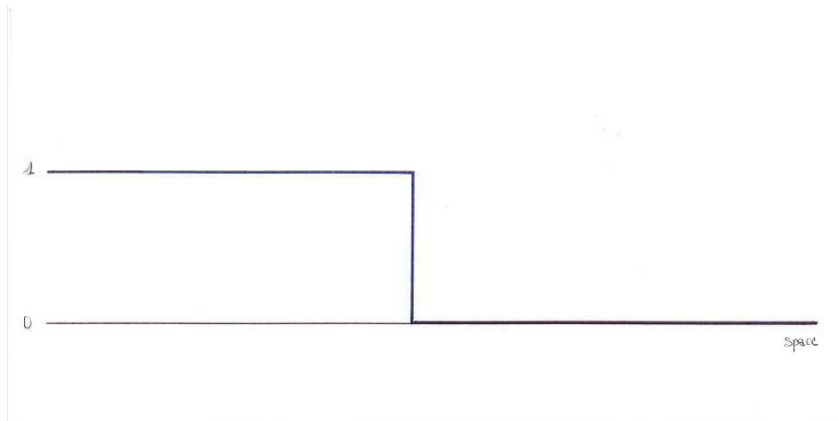
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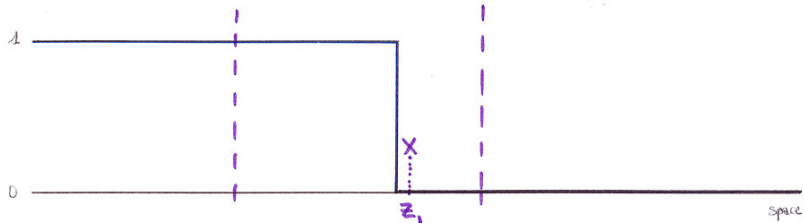
Very flexible framework: Gaussian replacements, possibility to add natural selection, recombination, inhomogeneous environment, ...

Back to the high-tech example



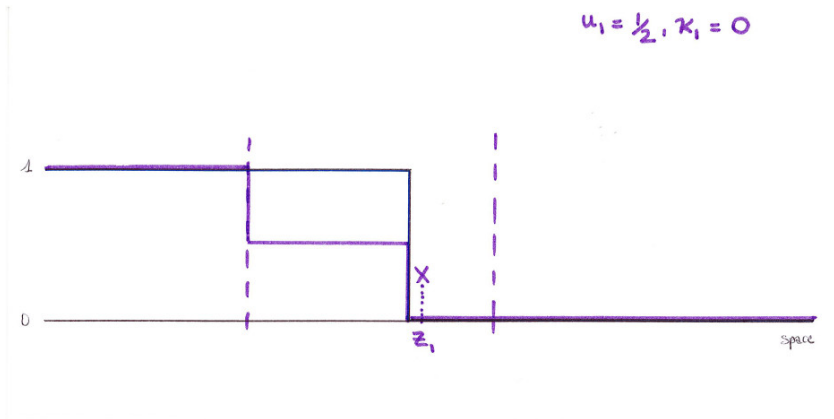
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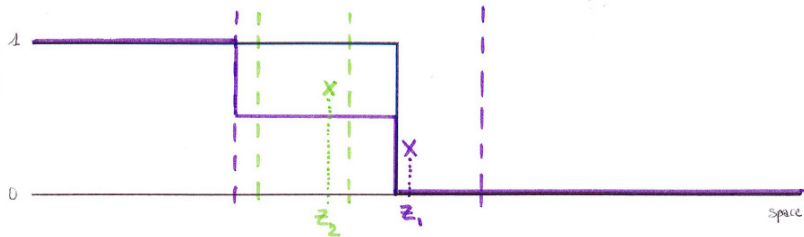
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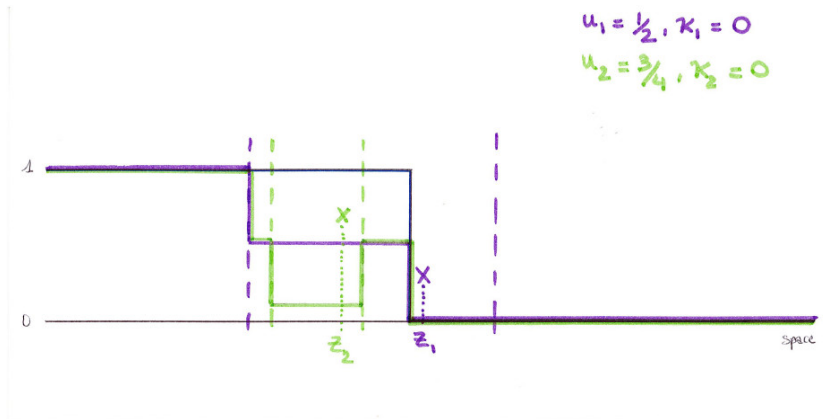
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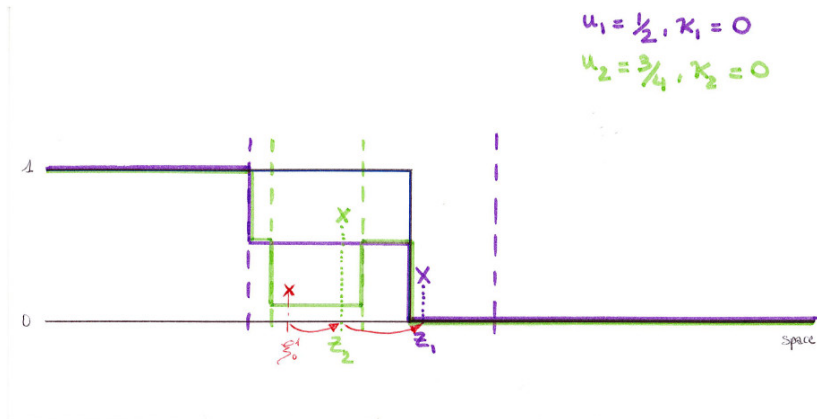
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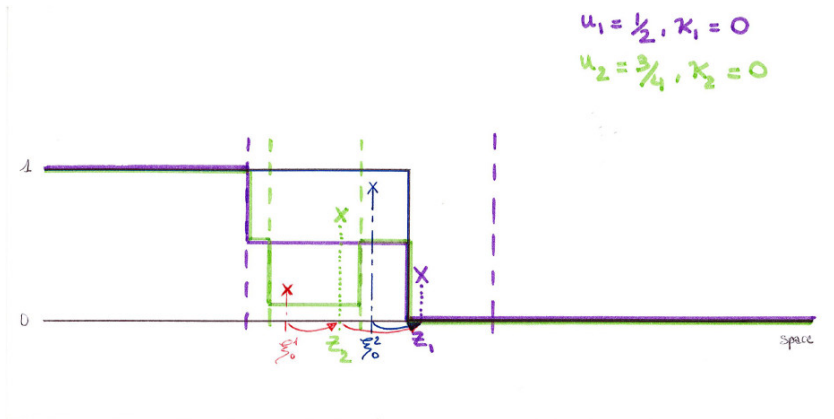
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Genealogies and duality

- ▶ We trace back the origin of the alleles observed in a sample of j individuals \rightsquigarrow “genealogical” process
- ▶ $(\{\xi_s^1, \dots, \xi_s^{N_s}\})_{s \geq 0}$ system of correlated symmetric jump processes, which coalesce/merge when they are *affected* by the same event.
- ▶ Jump rate of an ancestral lineage (presently at 0):

$$\int_{\mathbb{R}^d} \int_0^\infty \int_0^1 \mathbf{1}_{B(x,r)}(0) u \nu_r(du) \mu(dr) dx = \int_0^\infty \int_0^1 u V_r \nu_r(du) \mu(dr) < \infty.$$

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- ▶ If $K = \{0, 1\}$ and $w_t(x) := \rho_t(x, \{1\})$: for every $\psi \in C_c((\mathbb{R}^d)^j)$,

$$\begin{aligned} \mathbb{E}_{w_0} \left[\int_{(\mathbb{R}^d)^j} \psi(x_1, \dots, x_j) \left\{ \prod_{i=1}^j w_t(x_i) \right\} dx_1 \cdots dx_j \right] \\ = \int_{(\mathbb{R}^d)^j} \psi(x_1, \dots, x_j) \mathbf{E}_{\{x_1, \dots, x_j\}} \left[\prod_{i=1}^{N_t} w_0(\xi_t^i) \right] dx_1 \cdots dx_j. \end{aligned}$$

Long term evolution at an interface

(Berestycki, Etheridge & V., 2013)

Geographical space : \mathbb{R}^d **Allele space:** $\{0, 1\}$

► **Case 1: Constant radius**

We fix $R > 0$ and $u \in (0, 1]$. All events have radius R and impact u .

↪ Reproductions are purely local.

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▶ Case 2: Heavy-tailed radii

We fix $\alpha \in (1, 2)$ and $u \in (0, 1]$. Intensity measure on radii given by

$$\mu(dr) = \frac{\mathbf{1}_{\{r>1\}}}{r^{d+\alpha+1}} dr.$$

↪ Allows the occurrence of rare but very large events.

↪ Ancestral lineages behave like α -stable processes.

Time- and space-scales

- ▶ Case 1: Constant radius and impact
- ▶ Case 2: Constant impact, intensity $r^{-(d+\alpha+1)} dr$ for radii.

We set $\alpha = 2$ in case 1 and for every $n \geq 1$,

$$w_t^n(x) := w_{nt}(n^{1/\alpha}x).$$

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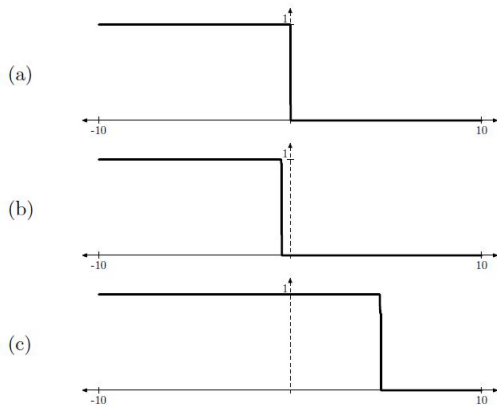
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Initial condition: $w_0(x) = \mathbf{1}_H(x)$, where $H = \{x_{(1)} \leq 0\}$.

Questions: How does w_t^n behave when n is large? Width of the interface? Resulting pattern of genetic diversity?

Purely local reproductions, $d = 1$



$u = 0.8$, $r = 0.033$ and $n = 10^3$. Initial condition, after 10^5 events, after 10^7 events.
(Simulations by J. Kelleher, Oxford Univ.)

That is...

Theorem 1 (Berestycki, Etheridge & V., 2013)

- ▶ There exists a process $(M_t^{(2)}, t \geq 0)$, with values in \mathcal{M}_λ , such that

$$M^n \xrightarrow{(f.d.d.)} M^{(2)} \quad \text{as } n \rightarrow \infty.$$

- ▶ Moreover, there exists $\tilde{\sigma}^2 > 0$ such that, if X denotes standard Brownian motion and

$$p_t^{(2)}(x) := \mathbf{P}_x[X_{U\tilde{\sigma}^2 t} \in H], \text{ then}$$

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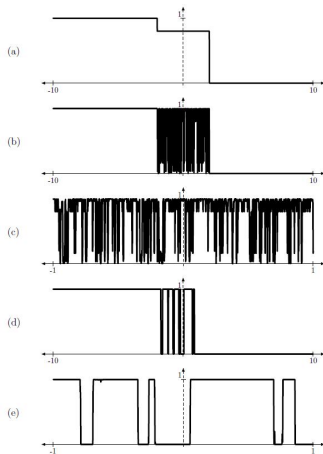
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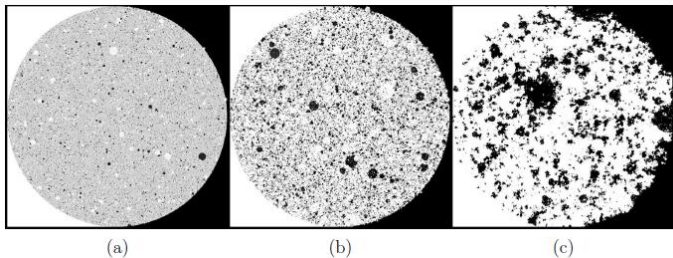
- ↪ If $d \geq 2$: for all $t \geq 0$, $w_t^{(2)}(x) = p_t^{(2)}(x)$ Lebesgue-a.e.

In the presence of catastrophes, $d = 1$



$u = 0.8$, $\alpha = 1.3$ and $n = 10^4$ (Simulations by J. Kelleher).
(a) Initial condition, (b-c) after 100 events, (d-e) after 10^6 events.

In the presence of catastrophes, $d = 2$



$u = 0.8$, $\alpha = 1.3$ and $n = 10^3$. After 10^5 , 10^6 and 10^7 events.
(Simulations by J. Kelleher)

Asymptotic behaviour in the presence of large events

Theorem 2 (Berestycki, Etheridge & V., 2013)

- ▶ There exists a process $(M_t^{(\alpha)}, t \geq 0)$, with values in \mathcal{M}_λ , such that

$$M^n \xrightarrow{(f.d.d.)} M^{(\alpha)} \quad \text{as } n \rightarrow \infty.$$

- ▶ Moreover, there exists a symmetric α -stable process $X^{(\alpha)}$ such that, if

$$p_t^{(\alpha)}(x) := \mathbf{P}_x[X_{ut}^{(\alpha)} \in H],$$

then *in any dimension*, for all $t > 0$, $w_t^{(\alpha)}$ is a random field of correlated Bernoulli r.v. satisfying

$$\mathbb{E}[w_t^{(\alpha)}(x)] = p_t^{(\alpha)}(x).$$

Conclusions

- ▶ **No local coexistence of alleles** except if $d \geq 2$ and reproductions are purely local.
- ▶ The correlations between the local allele frequencies are characterised by the genealogical process.

Correlation length:

- ▶ \sqrt{n} in the case of constant (or bounded) event radius,
- ▶ $n^{1/\alpha}$ in the case of heavy tailed radius distribution.

↪ Rare but massive extinction/recolonisation events may have a significant impact on the genetic diversity seen in a population.

An example of spatial clustering - *C. elegans* in the wild (M.-A. Félix & H. Teotonio - ENS)



- ▶ A complex life cycle, with an optional *dauer* phase during which it only moves - does not eat/reproduce.
- ▶ Local population dynamics in “boom and bust”:
 - ▶ Appearance of a food source \rightsquigarrow exponential growth
 - ▶ Exhaustion of the resource \rightsquigarrow transition to *dauer* stage
- ▶ Migration possible via individual motion (slow) and/or by hitchhiking in groups (~ 10 individuals, potentially moving several meters hung on a snail or an isopod).

Why you may doubt

- ▶ For a 1mm long nematode, a 10m wide strip is not particularly one-dimensional...
- ▶ All we can conclude is that a neutral model is compatible with this type of spatial segregation, but maybe (or almost surely...) individuals are in competition for resources, and subject to selection pressures due to pathogens.
 ↪ “Boom and bust” dynamics reinforcing a selective pattern?
- ▶ Still, in “2d” (apple orchard in Orsay), minor haplotypes are maintained and spatial segregation is not observed.

Reference: A. Richaud, G. Zhang, D. Lee, J. Lee, M.-A. Félix (2018). The local coexistence pattern of selfing genotypes in *Caenorhabditis elegans* natural metapopulations. *Genetics*, 208(2), 807–821.

The effects of a weak selection pressure

We bias the choice of the parent by giving a slight reproductive advantage to individuals with allele 0 over individuals with allele 1.

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► **One way to proceed:**

$\rightsquigarrow \Pi^N$ PPP on $\mathbb{R} \times \mathbb{R}^d \times \mathbb{R}_+^* \times [0, 1]$ with intensity $dt \otimes dx \otimes \mu(dr)\nu_r(du)$.
Local update of the allele frequencies: we draw κ according to the allelic distr. in the area of the event, then $\forall y \in B(x_i, r_i)$,

$$w_{t_i}(y) = (1 - u_i)w_{t_i-}(y) + u_i \mathbf{1}_{\{\kappa=1\}}.$$

$\rightsquigarrow \Pi^S$ PPP on $\mathbb{R} \times \mathbb{R}^d \times \mathbb{R}_+^* \times [0, 1]$, indep. of Π^N and with intensity $s' dt \otimes dx \otimes \mu(dr)\nu_r(du)$. This time, we choose two alleles κ, κ' indep. and according to the allelic distr. in the area of the event, and then $\forall y \in B(x_i, r_i)$,

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► **Dual process of *potential* ancestors**, where ancestral lineages can also branch during the events of Π^S .

Weak selection and high population density

(Etheridge, V. & Yu, 2020)

$$u_n = \frac{u}{n^\gamma}, \quad s_n = \frac{s}{n^\delta}, \quad w_t^n(x) = w_{nt}(n^\beta x),$$

or rather:

$$\bar{w}_t^n(x) := \frac{n^{d\beta}}{V_R} \int_{B(x, n^{-\beta} R)} w_t^n(y) dy \quad \text{and} \quad \bar{M}_t^n \text{ with 'density' } \bar{w}_t^n.$$

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Again, for every $n \in \mathbb{N}$

↪ **Case 1:** all events are of fixed radius $R > 0$ and impact u_n .

↪ **Case 2:** the intensity measure on radii is

$$\mu(dr) = \frac{\mathbf{1}_{\{r>1\}}}{r^{d+\alpha+1}} dr, \quad \alpha \in (1, 2),$$

and the impact u_n is fixed. (And we take $R = 1$ in the def. of $\bar{w}_t^n(x)$.)

Case of purely local reproductions

Theorem 3 (Etheridge, V. & Yu, 2020)

$$u_n = un^{-1/3}, \quad s_n = sn^{-2/3}, \quad w_t^n(x) = w_{nt}(n^{1/3}x)$$

If \bar{M}_0^n converges to $M_0 \in \mathcal{M}_\lambda$ as $n \rightarrow \infty$, then $(\bar{M}_t^n)_{t \geq 0}$ converges in law in $D_{\mathcal{M}_\lambda}[0, \infty)$ to a Markov process M^∞ , with initial value M_0 and characterised by: for all families $\{w_t^\infty, t \geq 0\}$ of representatives of the density of each M_t^∞ and every $f \in C_c^\infty(\mathbb{R}^d)$,

$$\left(\langle w_T^\infty, f \rangle - \langle w_0^\infty, f \rangle - \int_0^T \left\{ \frac{u\Gamma_R}{2} \langle w_t^\infty, \Delta f \rangle - V_R u s \langle w_t^\infty (1 - w_t^\infty), f \rangle \right\} dt \right)_{T \geq 0}$$

is a martingale with quadratic variation zero if $d \geq 2$ and

$$4R^2 u^2 \int_0^T \langle w_t^\infty (1 - w_t^\infty), f^2 \rangle dt$$

if $d = 1$, where $\Gamma_R > 0$ depends only on d and R .

In the presence of large events

Theorem 4 (Etheridge, V. & Yu, 2020)

$$u_n = un^{-\frac{\alpha-1}{2\alpha-1}}, \quad s_n = sn^{-\frac{\alpha}{2\alpha-1}}, \quad w_t^n(x) = w_{nt}(n^{\frac{1}{2\alpha-1}}x)$$

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Next questions

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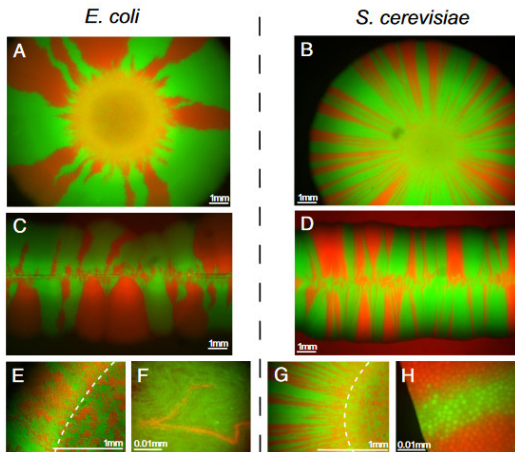
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- ▶ Impact of stochasticity on range expansions (Louvet 2023, **Louvet & V. 2023**).
↪ A very interesting question is that of *gene surfing* (cf. work by O. Hallatschek, D.R. Nelson, ...).

Genetic diversity in expanding populations



Hallatschek *et al.* (2007), PNAS. Sectors forming after placing a well-mixed droplet (or line) of fluorescent-green/red bacteria (left) or yeast (right).

Merci pour votre attention !